Third homework assignment PhD in Business Economics
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Advanced Bayesian Econometrics
Due at 9am, February 18th, 2021.

Use, preferably, Rmarkdown (via RStudio) to produce your report in PDF or HTML.

## Nonlinear Gaussian regression

Let us consider the context of Example 6.1 (pages 192-194) of Gamerman and Lopes (2006), where the response variable $y$, is the velocity of an enzymatic reaction (in counts $/ \mathrm{min} / \mathrm{min}$ ) and the regressor, $x$, is substrate concentration (in ppm). Check the book webpage at http://www.dme.ufrj.br/mcmc/ chapter6.html. Here is the data and a scatter plot showing the nonlinear relationship between $y$ and $x$ :

```
x = c(0.02,0.02,0.06,0.06,0.11,0.11,0.22,0.22,0.56,0.56,1.10,1.10)
y = c(76,47,97,107,123,139,159,152,191,201,207,200)
plot(x,y)
```

Model. We would like to entertain a Gaussian nonlinear model, for $i=1, \ldots, n(n=12)$,

$$
y_{i} \mid x_{i}, \beta, \gamma, \sigma^{2} \sim N\left(\beta_{0}+\beta_{1} g\left(x_{i}, \gamma\right), \sigma^{2}\right)
$$

where $\beta=\left(\beta_{0}, \beta_{1}\right)^{\prime}$ and $g\left(x_{i}, \gamma\right)=x_{i} /\left(\gamma+x_{i}\right)$, for $\beta \in \mathbb{R}^{2}, \gamma \in \mathbb{R}$ and $\sigma^{2} \in \mathbb{R}^{+}$.

Prior. Let us consider the prior for $\beta$ and $\sigma^{2}$ as follows:

$$
\begin{aligned}
p\left(\beta, \gamma, \sigma^{2}\right) & =p\left(\beta \mid \sigma^{2}\right) p(\gamma) p\left(\sigma^{2}\right) & & \\
\beta \mid \sigma^{2} & \sim N\left(b_{0}, \sigma^{2} B_{0}\right), & & b_{0}=(50,170)^{\prime}, B_{0}=3 I_{2} \\
\gamma & \sim N\left(g_{0}, \tau_{0}^{2}\right), & & g_{0}=0, \tau_{0}^{2}=1 \\
\sigma^{2} & \sim I G\left(\nu_{0} / 2, \nu_{0} \sigma_{0}^{2} / 2\right), & & \nu_{0}=5, \sigma_{0}^{2}=10 .
\end{aligned}
$$

Let $x=\left(x_{1}, \ldots, x_{n}\right)$ and $y=\left(y_{1}, \ldots, y_{n}\right)$. Answer the following questions.
(a) Show that $p\left(\beta \mid x, y, \sigma^{2}, \gamma\right)$ is a Gaussian distribution.
(b) Show that $p\left(\sigma^{2} \mid x, y, \beta, \gamma\right)$ is an Inverse-Gamma distribution.
(c) Show that $p\left(\sigma^{2} \mid x, y, \gamma\right)$ is also an Inverse-Gamma distribution.

Note 1: (c) is possible because we made the prior of $\beta$ conditional on $\sigma^{2}, p\left(\beta \mid \sigma^{2}\right)$, which results in $p\left(\sigma^{2} \mid x, y, \gamma\right)$ also being an Inverse-Gamma distribution. What is the catch? Well, the catch is that multiplying (a) and (c) is exactly $p\left(\beta, \sigma^{2} \mid x, y, \gamma\right)$, while iterating between (a) and (b) is approximately $p\left(\beta, \sigma^{2} \mid x, y, \gamma\right)$, which is the standard Gibbs (or, more generally, MCMC) theorem.

Note 2: It is not hard to see that $\left(\gamma \mid x, y, \beta, \sigma^{2}\right)$ comes from a distribution of no known form, since $\gamma$ from the prior does not conjugate with $\gamma$ from the likelihood function. Nonetheless, it can be point-wise evaluated up to a normalizing constant as

$$
\begin{aligned}
p\left(\gamma \mid x, y, \beta, \sigma^{2}\right) & \propto \exp \left\{-\frac{1}{2 \tau_{0}^{2}}\left(\gamma^{2}-2 \gamma m_{0}\right)\right\} \\
& \times \exp \left\{-\frac{1}{2 \sigma^{2}}\left[\beta_{1}^{2} \sum_{i=1}^{n} g^{2}\left(x_{i}, \gamma\right)-2 \sum_{i=1}^{n}\left(y_{i}-\beta_{0}\right) g\left(x_{i}, \gamma\right)\right]\right\} .
\end{aligned}
$$

(d) Algorithm 1: Implement an MCMC algorithm that cycles through the full conditionals:
(d1) Gibbs step: Sample $\sigma^{2}$ from $p\left(\sigma^{2} \mid x, y, \beta, \gamma\right)$;
(d2) Gibbs step: Sample $\beta$ from $p\left(\beta \mid x, y, \sigma^{2}, \gamma\right)$;
(d3) Metropolis-Hastings step: Sample Sample $\gamma$ from $p\left(\gamma \mid x, y, \beta, \sigma^{2}\right)$.
(e) Algorithm 2: Replace step (d1) from Algorithm 1 by
(e1) Gibbs step: Sample $\sigma^{2}$ from $p\left(\sigma^{2} \mid x, y, \gamma\right)$.
Conceptually, what is the difference between the above MCMC algorithms?
(f) Compare the algorithms in terms of MCMC mixing, sample autocorrelation functions and effective sample sizes, as well as by comparing the approximate marginal posterior distributions for $\beta_{0}, \beta_{1}, \gamma$ and $\sigma^{2}$.
(g) Finally, on the top of the scatterplot of $x$ against $y$, add the posterior predictive curve. More precisely, for a grid of new values of $x$, say $x_{n+1}$, in $\{0.02,0.03, \ldots, 1.10\}$, a 109-point grid, draw the 2.5 th, 50 th and 97.5 th percentiles the posterior predictive densities

$$
p\left(y_{n+1} \mid x_{n+1}, x, y\right)=\int p\left(y_{n+1} \mid x_{n+1}, \beta, \gamma, \sigma^{2}\right) p\left(\beta, \gamma, \sigma^{2} \mid x, y\right) d \beta d \gamma d \sigma^{2}
$$

Recall that, by Monte Carlo integration,

$$
\widehat{p}_{m c}\left(y_{n+1} \mid x_{n+1}, x, y\right)=\frac{1}{M} \sum_{i=1}^{M} p\left(y_{n+1} \mid x_{n+1}, \beta^{(i)}, \gamma^{(i)}, \sigma^{2(i)}\right),
$$

where $\left\{\left(\beta, \gamma, \sigma^{2}\right)^{(i)}\right\}_{i=1}^{M}$ are draws from the posterior $p\left(\beta, \gamma, \sigma^{2} \mid x, y\right)$, which could be obtained from algorithm 1 or 2 above. In fact, draws $\left\{\left(\beta, \gamma, \sigma^{2}\right)^{(i)}\right\}_{i=1}^{M}$ can also be used to generate draws $\left\{y_{n+1}^{(i)}\right\}_{i=1}^{M}$ from $p\left(y_{n+1} \mid x_{n+1}, x, y\right)$ by sampling $y_{n+1}^{(i)}$ from $p\left(y \mid x_{n+1},\left(\beta, \gamma, \sigma^{2}\right)^{(i)}\right)$, for $i=1, \ldots, M$.

