

GIBBS SAMPLER:

Poisson with a change point

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February 2014

Coal mining disasters in Great Britain

Annual counts of coal mining disasters from 1851 to 1962¹.

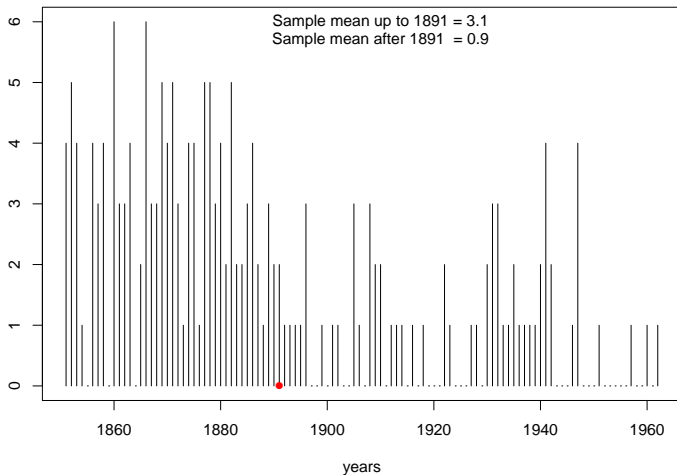
4	5	4	1	0	4	3	4	0	6
3	3	4	0	2	6	3	3	5	4
5	3	1	4	4	1	5	5	3	4
2	5	2	2	3	4	2	1	3	2
2	1	1	1	1	3	0	0	1	0
1	1	0	0	3	1	0	3	2	2
0	1	1	1	0	1	0	1	0	0
0	2	1	0	0	0	1	1	0	2
3	3	1	1	2	1	1	1	1	2
4	2	0	0	0	1	4	0	0	0
1	0	0	0	0	0	1	0	0	1
0	1								

Scientific question: There is a suspicion of a change point m along the observation process where the count means change, $m = 1, \dots, n$.

¹Jarret (1979) A note on the intervals between coal-mining disasters. *Biometrika*, **66**, 191-193.

Time series

Counts of coal mining disasters in Great Britain



Model²

Model: For a given change point m , the observations follow

$$y_t | \lambda, m \sim \text{Poi}(\lambda) \quad t = 1, \dots, m$$

$$y_t | \phi, m \sim \text{Poi}(\phi) \quad t = m + 1, \dots, n.$$

Recall that $X \sim \text{Poi}(\theta)$ when

$$\Pr(X = k | \theta) = \frac{e^{-\theta} \theta^k}{k!} \quad k = 0, 1, 2, \dots,$$

with $E(X | \theta) = \theta$.

²Carlin, Gelfand and Smith (1992) Hierarchical Bayesian analysis of change-point problems. *Applied Statistics*, **41**, 389-405. Tanner (1996, page 147) *Tools for Statistical Inference: Methods for the Exploration of Posterior Distributions and Likelihood Functions*, 3rd edition, Springer-Verlag: New York.

Prior specification

The model is completed with independent prior distributions $\lambda \sim G(\alpha, \beta)$, $\phi \sim G(\gamma, \delta)$ and m uniformly distributed over $\{1, \dots, n\}$ where α, β, γ and δ are known constants.

More precisely,

$$p(\lambda|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda} \quad \text{for } \lambda, \alpha, \beta > 0$$

$$p(\phi|\gamma, \delta) = \frac{\delta^\gamma}{\Gamma(\gamma)} \phi^{\gamma-1} e^{-\delta\phi} \quad \text{for } \phi, \gamma, \delta > 0$$

$$Pr(m) = \frac{1}{n} \quad \text{for } m = 1, \dots, n.$$

Posterior

$$\begin{aligned} p(\lambda, \phi, m|y) &\propto \left(e^{-m\lambda} \lambda^{\sum_{t=1}^m y_t} \right) \left(e^{-(n-m)\phi} \phi^{\sum_{t=m+1}^n y_t} \right) \\ &\times \left(\lambda^{\alpha-1} e^{-\beta\lambda} \right) \left(\phi^{\gamma-1} e^{-\delta\phi} \right) \\ &\propto \left(\lambda^{(\alpha+s_m)-1} e^{-(\beta+m)\lambda} \right) \left(\phi^{(\gamma+s_n-s_m)-1} e^{-(\delta+n-m)\phi} \right) \end{aligned}$$

where $s_m = \sum_{t=1}^m y_t$ for $m = 1, \dots, n$.

The full conditional densities of λ and ϕ are

$$(\lambda|m, y) = G(\alpha + s_m, \beta + m)$$

$$(\phi|m, y) = G(\gamma + s_n - s_m, \delta + n - m)$$

while

$$Pr(m|\lambda, \phi, y) \propto \lambda^{\alpha+s_m-1} e^{-(\beta+m)\lambda} \phi^{\gamma+s_n-s_m-1} e^{-(\delta+n-m)\phi}$$

for $m = 1, \dots, n$.

Exact calculations

It is possible to obtain $Pr(m|y)$ analytically:

$$Pr(m|y) \propto \frac{\Gamma(\alpha + s_m)\Gamma(\gamma + s_n - s_m)}{(m + \beta)^{\alpha + s_m}(n - m + \delta)^{\gamma + s_n - s_m}},$$

for $m = 1, \dots, n$, Therefore,

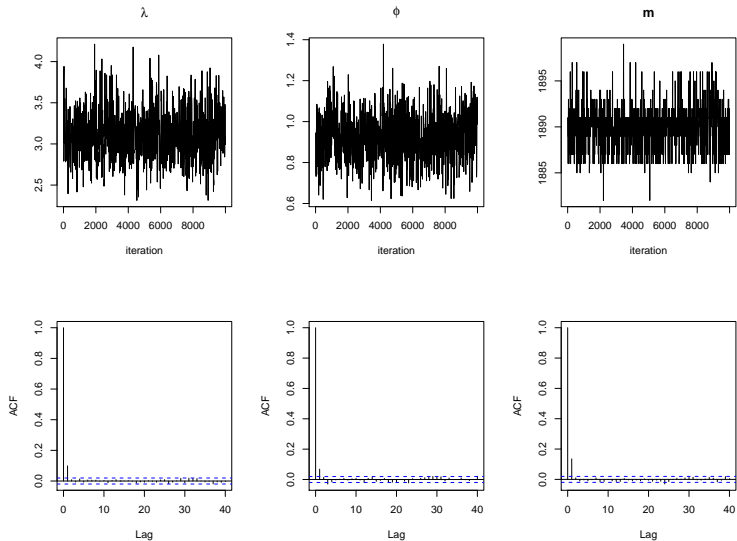
$$p(\lambda|y) = \sum_{m=1}^n p(\lambda|y, m)Pr(m|y)$$

$$p(\phi|y) = \sum_{m=1}^n p(\phi|y, m)Pr(m|y)$$

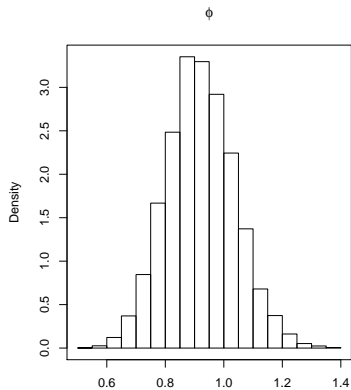
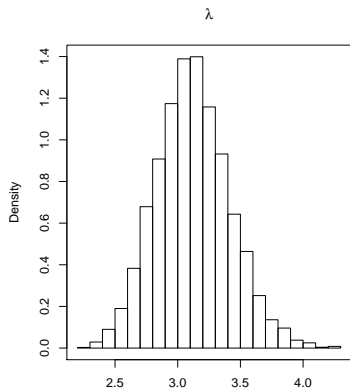
For instance

$$E(\lambda|y) = \sum_{m=1}^n E(\lambda|m, y)Pr(m|y)$$

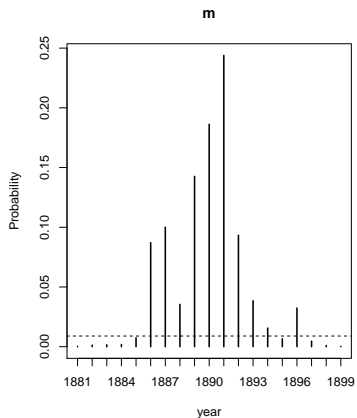
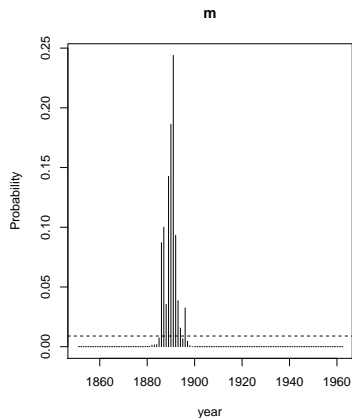
Coal mining disasters: MCMC output



Coal mining disasters: Posterior inference of λ and ϕ



Coal mining disasters: Posterior inference of m



Dashed line is the prior $Pr(m) = 1/n \quad \forall m$.

Coal mining disasters: Posterior summaries

Gibbs approximations are based on 10,000 draws, starting at $m^{(0)} = 1891$ and after discarding the first 10,000 draws.

Par.	Exact			Gibbs sampler		
	Mean	St.Dev.	95% C.I.	Mean	St.Dev.	95% C.I.
λ	3.12	0.28	(2.57,3.72)	3.12	0.29	(2.58,3.73)
ϕ	0.92	0.11	(0.68,0.96)	0.92	0.12	(0.70,1.16)
m	1890	2.42	(1886,1895)	1890	2.42	(1886,1896)

C.I. stands for credibility interval.